

Hyaluronic acid/ Hyaluronidase as biomarkers for bladder cancer: a diagnostic meta-analysis

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This study aimed to determine the value of HA/HAase for detecting bladder cancer on the basis of preceding statistical performance. PubMed, Springer Link, Web of Science and Cochrane Library were systematically searched to identify potentially relevant published articles by using the key words: “bladder cancer or bladder tumor or bladder carcinoma”, “hyaluronic acid or hyaluronan”, “hyaluronidase or HAase”. The methodological quality of each study was assessed by QUADAS-2. According to the inclusive and exclusive criteria, 8 articles were identified and methodologically analyzed by STATA 12.0 software package. The results showed that the pooled sensitivity of HA and HAase was 0.832 (95% confidence interval [CI]: 0.798, 0.861) and 0.834 (95% CI: 0.756, 0.891) respectively, the pooled specificity was 0.886 (95% CI: 0.852, 0.913) and 0.860 (95% CI: 0.801, 0.904), and the area under the summary ROC curve (AUC) was 0.90 (95% CI: 0.87, 0.92) and 0.91 (95% CI: 0.88, 0.93), respectively. Simultaneously the diagnostic accuracy of the combination of HA and HAase showed that the pooled sensitivity was 0.908 (95% CI: 0.879, 0.931), the pooled specificity was 0.825 (95% CI: 0.789, 0.856) and AUC was 0.94 (95% CI: 0.91, 0.95), indicating a relatively higher accuracy than HA and HAase alone. This meta-analysis strongly suggests that HA/HAase could be used as biomarkers for the diagnosis of bladder cancer.

Key words: bladder cancer, hyaluronic acid, hyaluronidase, diagnosis, meta-analysis

Bladder cancer (BC), the most common carcinoma in urinary system [1], is the twelfth most frequently diagnosed cancer all over the world, and the sixth in developed country and sixteenth in developing country [2]. The majority of bladder cancers occur in males and there is nearly a 14-fold variety in incidence among different countries [3]. In general, only when a patient presents with hematuria, bladder cancer could be then detected [4]. Cytology and cystoscopy are the principle methods for the diagnosis of patients with bladder cancers [5]. However, these two kinds of approaches have low sensitivity and specificity, consequently, early diagnosis of bladder cancer remains a challenge [6]. Meanwhile, despite significant advances in the molecular pathology of bladder cancer, it is still a significant health problem [7]. Given the importance of early diagnosis for bladder cancer, new biomarkers to reduce the frequency of BC are conceivable [8]. Hyaluronidase (HAase) is a kind of glycosidase which mainly degrade hyaluronic acid (HA) [9], and HA is a glycosaminoglycan that has osmotic, homeostatic, and structural properties

in normal tissues [10]. Both HA and HAase are known to play important roles during embryonic development, vascular remodeling, immune surveillance, and tumor progression [11, 12]. Studies reported that HA regulates cell adhesion [13, 14], has an influence on cell proliferation and migration [15, 16], and the accumulation of HA in tumour interstitial fluid correlates with lymph node metastasis [17]. Lokeshwar VB' study showed that HAase was involved in tumor growth and tumor angiogenesis [18, 19]. Previous studies also reported that HA and HAase were both associated with several different kinds of carcinomas [20]. And the association between HA/HAase and bladder cancer has also been studied for a few years, but so far, there was no meta-analysis on the diagnostic value of HA/HAase as biomarkers for bladder cancer detection. What's more, the individual studies showed different diagnostic accuracy, for example, Lokeshwar and colleagues reported that urinary HA measurement has a sensitivity and specificity of 91.9% and 92.8% to detect bladder cancer, respectively [21], but in a cross-sectional study of 194 urine specimens (97

bladder cancer patients and 97 control individuals), the HA test showed a sensitivity of 82.5% and a specificity of 89.7%, in addition, the HA-HAase test showed 89.7% sensitivity and 83.5% specificity for bladder cancer [22]. Measurement of the levels of HA/HAase seems to be a highly accurate method for detecting bladder cancer. In this study, we are aiming to summarize the experimental studies to confirm the potential value of HA/HAase as BC marker. To our knowledge, this is the first meta-analysis for the assessment of the roles of HA/HAase in the diagnosis of bladder cancer.

Materials and methods

Search strategy, selection criteria. To identify all primary research literatures in which the value of hyaluronidase in BC were analyzed, electronic databases, including PubMed, Springer Link, Web of Science and Cochrane Library were used by searching key words: “bladder cancer or bladder tumor or bladder carcinoma”, “hyaluronic acid or hyaluronan or HA”, “hyaluronidase or HAase”. We only collected data from papers published in English, ruling out meeting or conference abstracts.

The included studies meet the following criteria: (1) original study; (2) HA/HAase serves as biomarkers for BC diagnosis; (3) the diagnosis of bladder cancer was based on cytology and cystoscopy; (4) sufficient information to build 2×2 tables for calculating sensitivity and specificity; (5) when multiple publications reported on the same or overlapping data, we used the most recent or largest population; (6) the publication language was confined to English.

The exclusive criteria are as follows: (1) studies from conference abstracts, letters, editorials or reviews; (2) studies without control groups; (3) studies without valid data (sensitivity, specificity, true positive, false positive, false negative and true negative).

Data extraction. Data were extracted from each study by two reviewers independently using pre-specified selection criteria. Decisions were made and discordances about study

selection were resolved by consensus or by involving a third assessor. The following information was extracted from the studies: first author, publication data, sample size, test method, true positive (TP), false positive (FP), false negative (FN) and true negative (TN).

Quality assessment. We carried out quality assessment by using the Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2) tool [23], which is composed of 14 questions [24]. The assessment is made up of four primary domains, including patient selection, index detection, reference norm, and flow and timing. Each domain is evaluated according to the risk of bias, and the first three domains are assessed on the basis of applicability. Each item is marked as “high”, “low” or “unclear”, matches to high risk, low risk, and unclear, respectively.

Statistical analysis. The accuracy indicators for this diagnostic meta-analysis includes the pooled sensitivity (SEN), pooled specificity (SPE), diagnostic odds ratio (DOR), positive likelihood ratio (PLR), negative likelihood ratio (NLR), and their 95% confidence interval (CI). The analysis of diagnostic accuracy pursuant to a Summary Receiver Operating Characteristic (SROC) curve and the area under curve (AUC) of the SROC. The I^2 and Q test were performed to assess heterogeneity. A value larger than 50% for I^2 or a P -value of less than 0.1 for Q test indicates significant heterogeneity, and correspondingly, the random-effect model should be applied. Deeks' funnel plot asymmetry test was used to check the publication bias. Meanwhile, Spearman correlation analysis was used to check the threshold effect. All statistical analyses were performed by using STATA 12.0.

Results

Study selection and description. Using the established search strategy, we totally found 440 potentially relevant papers (PubMed: 65; Springer: 276; Web of Science Science: 95; Cochrane Library: 6) (Figure 1). According to the inclusive criteria and exclusive criteria, 8 eligible studies were finally selected [10, 22, 25-30], as shown in Table 1. These studies were published between 2000 and 2014. Data were obtained from studies directly or extracted indirectly by calculating the number of people and relevant ratios.

Quality assessment of studies. Based on the QUADAS-2, the outcomes are shown in Table 1. The QUADAS-2 scores of the enrolled studies in this article indicated that the overall quality of these included studies are generally good.

Diagnostic accuracy. The forest plot of SEN, SPE for HA/HAase assays is shown in Figure 2. The pooled SEN of the included studies of HA test was 0.832 (95% CI: 0.798, 0.861), the pooled SPE, PLR, NLR, DOR AUC and their 95% confidence interval were as following: 0.886 (95% CI: 0.852, 0.913), 7.290 (95% CI: 5.581, 9.523), 0.190 (95% CI: 0.157, 0.230), 38.335 (95% CI: 26.250, 55.9985) and 0.90 (95% CI: 0.87, 0.92), (shown in Table 2). The diagnostic accuracy of the overall results of HAase test and their 95% CI were as following: the SEN was 0.834 (95% CI, 0.756, 0.891), the SPE was 0.860 (95%

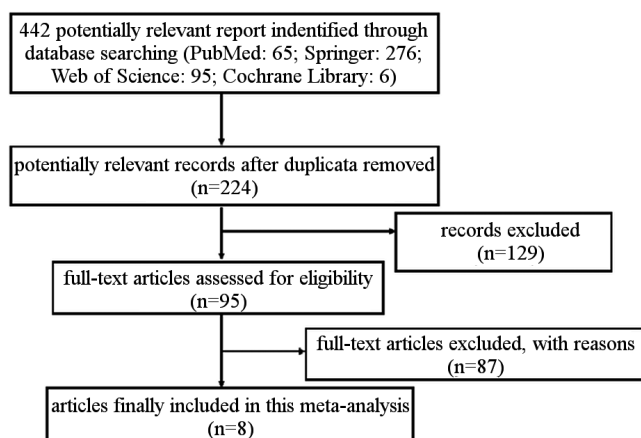


Figure 1. Flow diagram for selection of studies for the meta-analysis.

Table. 1. The characteristics of eligible studies for HA/ HAase test

| No | First Author | Year | Country | QUADAS-2 | Sampling object | Sample size | TP | FP | FN | TN |
|----|----------------------------|------|---------|----------|-----------------|-------------|-----|----|----|-----|
| 1 | Lokeshwar ^[10] | 2000 | USA | 10 | HA | 504 | 217 | 24 | 44 | 219 |
| | | | | | HAase | 504 | 159 | 50 | 36 | 259 |
| | | | | | HA-HAase | 504 | 240 | 38 | 21 | 205 |
| 2 | Hautmann ^[22] | 2001 | USA | 10 | HA | 46 | 31 | 0 | 3 | 12 |
| | | | | | HAase | 51 | 27 | 4 | 4 | 16 |
| | | | | | HA-HAase | 51 | 30 | 5 | 1 | 15 |
| 3 | Jamshidian ^[25] | 2011 | Iran | 12 | HA | 194 | 80 | 10 | 17 | 87 |
| | | | | | HAase | 194 | 60 | 23 | 8 | 103 |
| | | | | | HA-HAase | 194 | 87 | 16 | 10 | 81 |
| 4 | Passerotti ^[26] | 2011 | Brazil | 10 | HA | 350 | 131 | 37 | 29 | 153 |
| 5 | Eissa ^[27] | 2012 | Egypt | 12 | HAase | 216 | 89 | 11 | 11 | 105 |
| 6 | Nossier ^[28] | 2014 | Egypt | 10 | HAase | 66 | 26 | 1 | 14 | 25 |
| 7 | Schroeder ^[29] | 2004 | Germany | 10 | HA-HAase | 138 | 52 | 15 | 7 | 64 |
| 8 | Hautmann ^[30] | 2004 | USA | 12 | HA-HAase | 94 | 25 | 14 | 5 | 50 |

* ELISA-like assay was used for HA/HAase detection in all the included studies.

* TP: true positive; FP: false positive; FN: false negative; TN: true negative.

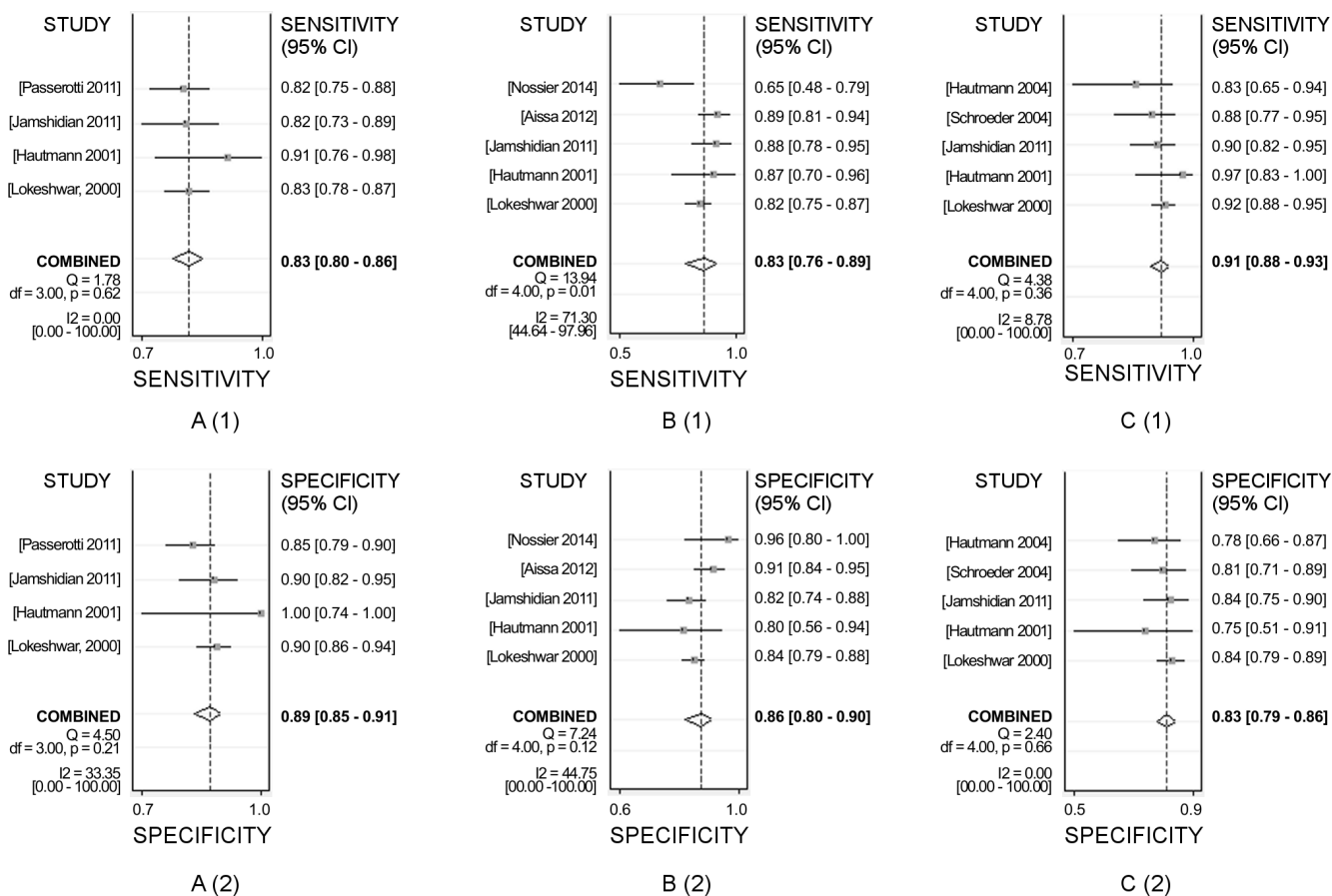


Figure 2. Forest plots of A(1) sensitivity for HA test in bladder cancer; A(2) specificity for HA test in bladder cancer; B(1) sensitivity for HAase test in bladder cancer; B(2) specificity for HAase test in bladder cancer; C(1) sensitivity for HA-HAase test in bladder cancer; C(2) specificity for HA-HAase test in bladder cancer.

Table 2. Sensitivity, specificity and accuracy calculations of the included HA, HAase and HA-HAase tests

| Category | HA test | HAase test | HA-HAase test |
|-------------|---------------------|---------------------|---------------------|
| Sensitivity | 0.832(0.798, 0.861) | 0.834(0.756, 0.891) | 0.908(0.879, 0.931) |
| Specificity | 0.886(0.852, 0.913) | 0.860(0.801, 0.904) | 0.825(0.789, 0.856) |
| AUC | 0.90(0.87, 0.92) | 0.91 (0.88, 0.93) | 0.94 (0.91- 0.95) |

CI: 0.801, 0.904), the PLR was 5.963 (95% CI: 4.215, 8.438), the NLR was 0.193 (95% CI: 0.131, 0.285), and the DOR was 30.869 (95% CI: 18.071, 52.730). The AUC of HAase tests was 0.91 (95% CI: 0.88, 0.93), indicating a relatively high diagnostic accuracy. The SROC curves for the included studies were shown in Figure 3. In addition, we also analyzed the pooled parameter of HA-HAase test, the AUC of HA-HAase was 0.94 (95% CI: 0.91, 0.95), showed a relatively higher diagnostic value than HA test and HAase test.

Threshold effect and heterogeneity analysis. During the research process, different studies could lead to different sensitivity and specificity, resulting in threshold effect and diagnostic odds ratios (DOR). If a threshold effect exists, the worker receiver operating characteristic (ROC) plane scatter distribution is in

a typical “shoulder arm-shaped” style, and the sensitivity would show a negative correlation to the specificity. In our analysis, the HA/HAase of SROC curve did not show this typical style (Figure 3), suggesting that there was no threshold effect.

In this paper, the diagnostic odds ratio was also used to explore the heterogeneity caused by non-threshold effect. For the HAase test, the results showed that the I^2 is 51.7%, indicating high heterogeneity among enrolled studies. To evaluate potential sources of between-study heterogeneity, we performed a further analysis. According to the original information, we studied tumor data by different grades (G1, G2 and G3). But unfortunately, limited by the samples in our meta-analysis, we got the tumor grades only from 3 articles (Table 3). And we found that the sensitivity of HAase for the diagnosis of G1 was much lower than G2 and G3, with sensitivity ranging from 81.2% to 100%.

Publication bias. Deeks’s funnel plot asymmetry test was used to evaluate publication bias. The P -value of the HA test, HAase test and HA-HAase test were 0.343, 0.563, 0.414, respectively. The results suggested that no significant publication bias existed among the included HA/HAase studies.

Clinical utility of index test. Fagan’s nomogram was shown for figuring up post-test probabilities (PTPs) (Figure 4). The Fagan’s nomogram for the included HA tests showed that when

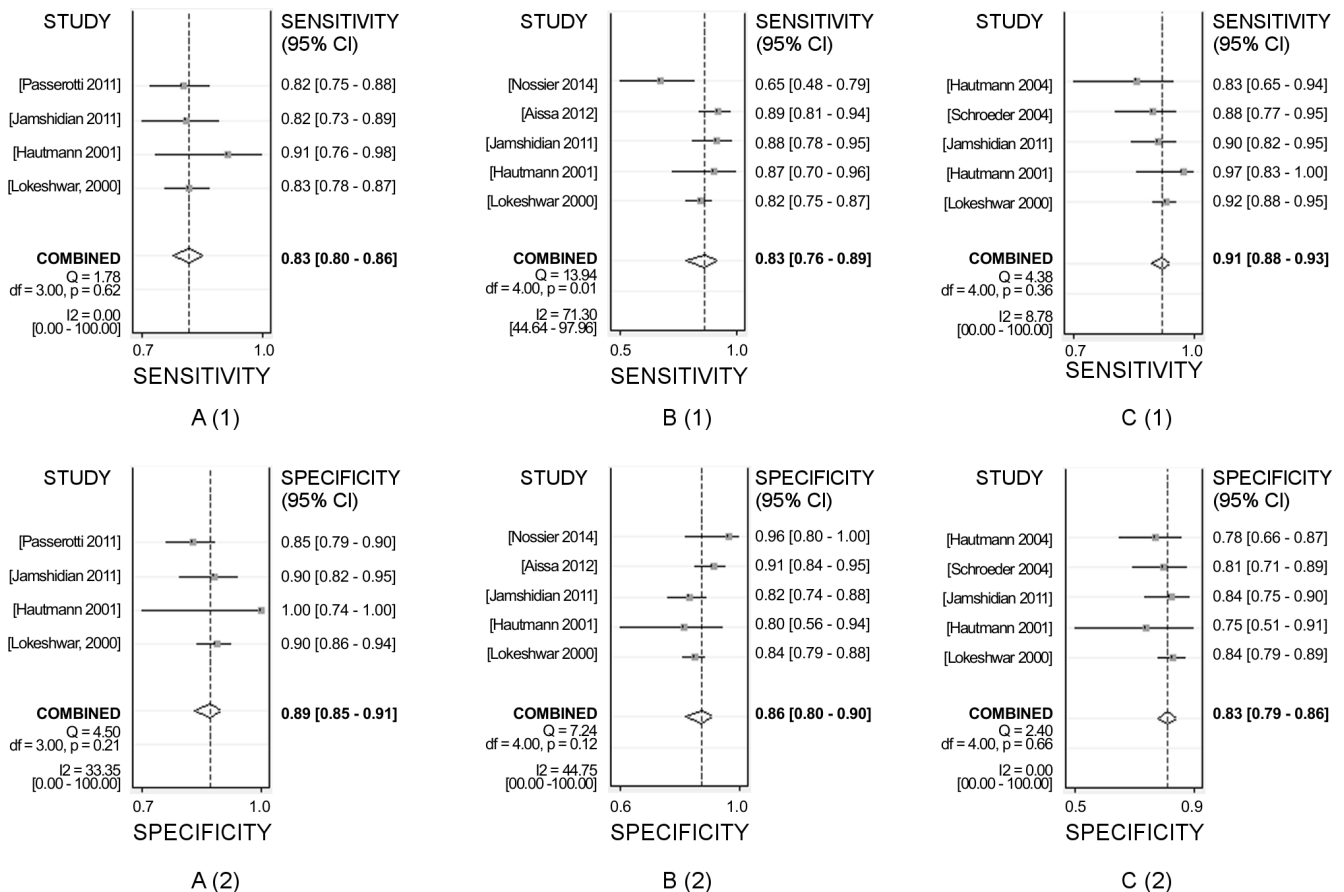


Figure 3. Summary ROC curve of HA (A), HAase (B), HA-HAase (C) test for pooled sensitivity and pooled specificity.

Table 3. The information of tumor grades for the included HAase tests.

| Author | year | Collecting time | Patients age (years) | Grade 1 | Grade 2 | Grade 3 |
|------------|------|-----------------|----------------------|---------|---------|---------|
| Lokeshwar | 2000 | 1995.1-1998.9 | 34-90 | 15/66 | 38/46 | 121/149 |
| Eissa | 2012 | NG | 37-78 | 65/76 | | 24/24 |
| Jamshidian | 2011 | 2007.7-2008.3 | 34-91 | 7/29 | 22/25 | 38/43 |

NG: not given

20% was used as the pre-test probability, the post-test probability would raise to 65% with a positive likelihood ratio of 7, and the probability would decrease to 5%, and the negative likelihood ratio is 0.19. For the HAase test, from the Fagan's nomogram, we found that when 20% was used as the pre-test probability, the post-test probability would raise to 60% with a positive likelihood ratio of 6, and the probability would decrease to 5%, and the negative likelihood ratio is 0.19. As for HA-HAase test, the Fagan's nomogram showed that when 20% was used as the pre-test probability, the post-test probability would rise to 56% with a positive likelihood ratio of 5, and the probability would decrease to 3%, and the negative likelihood ratio is 0.11.

Discussion

In the clinical practice, cystoscopy and urinary cytology are the most common methods for BC diagnosis [31]. Cystoscopy is the gold standard and urine cytology is the widely-used method. But the two methods also have shortcomings [32], they are costly, unpleasant, time consuming and trained personnel requirement [33]. At present, it is still a hot point to find

biomarkers for the diagnosis of bladder cancer [34, 35]. HA is known to promote tumor metastasis and help avoid immune surveillance by forming a protective coat around the tumor cells [36]. HAase, an endoglycosidase, degrades HA into small fragments that promote angiogenesis. Although HA/HAase has already been widely regarded as a biomarker [37-39], the small sample sizes and the lower statistical power of those single experiment limit their application. In this study, 8 papers with 1645 cases and controls were enrolled, the value of HA/HAase in bladder cancer diagnose was confirmed.

After a careful and serious data collection and analysis by software, we found that the pooled SEN of the included HA tests was 0.832, the pooled SPE 0.886. Both the sensitivity and specificity are greater than 0.80, indicating a good diagnostic accuracy. As for the HAase test, the pooled SEN was 0.834, the pooled SPE was 0.860, also showing a high potential diagnostic value of HAase as a noninvasive test, what's more, the pooled DOR was 30.869, indicating that the overall accuracy of HAase test for bladder cancer diagnosis is credible. As for SROC curve, the area under the summary ROC curve (AUC) of HA test was 0.90 and the AUC of HAase was 0.91. SROC is

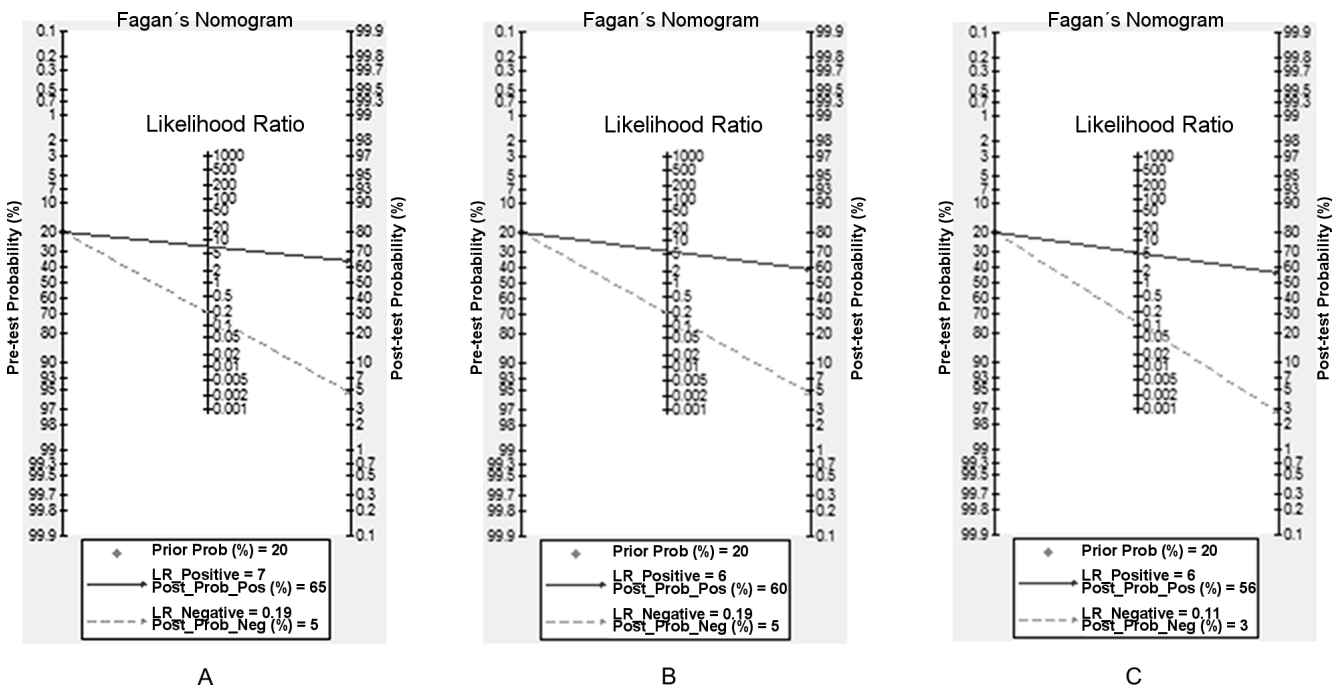


Figure 4 Fagan's nomogram: A straight line was used to contact the pretest probability of bladder tumor with HA, HAase, HA-HAase, by crossing the likelihood ratio line at a point that describes the results acquired.

normally used to sum up overall test performance and AUC serves as a measurement indicator [40]. In this article, the results showed that both HA test and HAase test had a good accuracy in diagnosing and testing BC with an AUC greater or equal to 0.90. In addition, the SROC curve, DOR and AUC values are not the only viable strategies for clinical diagnosis, the likelihood ratios (LR), including positive likelihood (PLR) and negative likelihood (NLR), also play a similar role in the evaluation of diagnostic accuracy [34]. For the HAase test, the pooled PLR of 5.963 suggested that patients with bladder cancer had a ~5.963-fold higher chance to have a positive result compared with controls. A pooled NLR of 0.193 meant that the probability of the individuals with bladder cancer was 19.3% when the HAase test was negative. All the results indicated that as a biomarker, HAase would be an accurate method for BC diagnosis. In this article, we compared the diagnostic value of HA test and HAase test with HA-HAase test, and found that the sensitivity of the combined test was 0.908, the pooled specificity was 0.825 and the AUC value was slightly increases to 0.94, indicating that the combination of HA and HAase had a better potential to be used as a biomarker for bladder cancer detection.

As mentioned above, different studies about the HA/HAase diagnostic value showed different results, and that was also the reason why we conducted this meta-analysis. From our study, the pooled sensitivity of HA test, HAase test and HA-HAase test were all greater than 0.80, even reached up to 0.904 for the combined test. The pooled specificity of the three methods were also greater than 0.80. These results indicated that HA/HAase had high sensitivity and specificity for early bladder cancer detection, which could make up the low sensitivity and specificity of cytology and cystoscopy. At present, several new biomarkers such as TERT and FGFR3 mutations, microRNA, telomerase, CD44 variants, cytokeratin 20 and others have shown their potential to be used as early diagnostic methods for detecting bladder cancer [41-44]. However, none of these markers are either as sensitive or as specific as the HA/HAase test. What's more, in this study, the AUC of the HA/HAase tests are all greater or equal to 0.90, strongly suggested that HA/HAase could be used as biomarkers for the diagnosis of bladder cancer.

Heterogeneity is a latent problem for the results of meta-analysis [45, 46], as non-homogeneous data are easy to cause misleading results. Among the different kinds of bias, the threshold effect must be considered firstly, so we used Spearman correlation coefficient to test the threshold effect, and its value was 0.029 ($P = 0.957$), suggesting that there was no heterogeneity from threshold effects. Previous studies showed that the HA/HAase levels gradually raised as the tumor grades increased. Lokeshwar and colleagues reported that the mean urinary HAase levels were 2.6- 7-fold higher in patients with G2 and G3 bladder cancer as compared with the levels in patients with G1 bladder cancer [10]. In this meta-analysis, the data showed that the I^2 for sensitivity of HAase test was 70.83% and the I^2 for specificity was 55.35%. As we are clear that sensitivity might be influenced by different tumor grades, then we performed a further analysis to determine the sensitiv-

ity of the HAase for detecting different grade bladder cancers. But, unfortunately, limited by the samples in our meta-analysis, we got the information of tumor grades from only 3 articles, and the data was not enough to conduct a subgroup analysis. Through a simple calculation, we found that the sensitivity of HAase in bladder cancer with G1 is much lower than those with G2/G3. And the sensitivity of HAase in bladder cancer patients with G2 or G3 tumors ranges from 81.2% to 100%, indicating that HAase has a good diagnostic value for G2 and G3 bladder cancer but not for G1 bladder cancer.

However, this meta-analysis still has several limitations. Firstly, only those papers written in English were included, relevant articles using other languages had not be enrolled in. Secondly, all records we searched are published articles, as a result, unpublished data were not included, some useful information may be lost with much possible. Although, Deeks' funnel plot asymmetry test showed that no significant publication bias among the included studies, however, the bias caused by those unpublished data should not be completely ignored. Thirdly, despite of the strict inclusive criteria, significant heterogeneity was still detected among the enrolled studies. To explore the source of the heterogeneity, we should have performed a subgroup analysis, but we could not effectively implemented because of the limited data. What's more, in clinical practice, bladder cancer is often treated by transurethral resection (TUR) or bacille Calmette-Gue´rin therapy (BCG) which could lead to an increase of urinary tract infection [47-50], in addition, HA is known to play an important role in the tissue response to injury and inflammation [51], and Bollyky PL also reported that injury and inflammation could enhance HA production [52]. The above-mentioned results suggested TUR or BCG might have effects on HA level. To figure out the effects of TUR or BCG, we analyzed the included papers, Passerotti [26], Nossier [28] and Schroeder [29] clearly pointed out that the samples they detected were all collected prior to treatment or surgery; but all of the studies conducted by Hautmanm [22], Jamshidian [25], Eissa [27] and Hautmanm [30] did not clarify whether the sample were collected prior to treatment or not; only in Lokeshwar's study^[10], they reported that among all the 261 bladder cancer patients, 72% were new cases and 28% were recurrences, and the patients with recurrence had been treated previously either by transurethral resection of the bladder (TUR) alone or followed by intra vesicle therapy (BCG or mitomycin C), and they pointed out that these treatments did not appear to affect the results. To further rule out the confounding, we need to develop a study to confirm whether TUR/BCG influence HA/HAase level or not. But we could not get an exact conclusion from current papers. In the future, researchers should collect samples from bladder cancer patients treated by TUR/BCG and matched controls, and then detect their HA/HAase level respectively, and finally, conduct a laboratory validation.

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